

Combining Covariate Adjustment with Group Sequential and Information Adaptive Designs to Improve Randomized Trial Efficiency

Outline

- 1 Background
- 2 Proposal: Combining Covariate Adjustment and GSDs
- 3 Proposal: Combining Covariate Adjustment and Information-Adaptive Designs
- 4 Simulation Study
- 5 Discussion

Covariate Adjustment

- In many clinical trials, data is collected on different **patient characteristics** at the time of entry
 - e.g., age, baseline severity and comorbidities

Covariate Adjustment

- In many clinical trials, data is collected on different **patient characteristics** at the time of entry
 - e.g., age, baseline severity and comorbidities
- **Covariate adjustment** is a statistical analysis method with high potential to **improve precision** for many of these trials.
 - **Pre-planned** adjustment for baseline variables when estimating **average treatment effect**.
 - Estimand is same as when using unadjusted estimator (e.g., difference in means).
 - **Goal**: avoid making any model assumptions beyond what's assumed for unadjusted estimator (**robustness to model misspecification**).

(e.g., Koch et al., 1998; Yang and Tsiatis, 2001; Rubin and van der Laan, 2008; Tsiatis et al., 2008; Moore and van der Laan, 2009b,a; Zhang, 2015; Jiang et al., 2018; Benkeser et al., 2020)

FDA Guidance: Example

- Primary endpoint Y : binary.

FDA Guidance: Example

- Primary endpoint Y : binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.

FDA Guidance: Example

- Primary endpoint Y : binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: G-computation/Standardization

- 1 Fit logistic regression model for

$$P(Y = 1|A, W) = \text{logit}^{-1}(\gamma_0 + \gamma_1 A + \gamma_2 W).$$

- 2 Compute standardized estimators for treatment specific means

- $\hat{E}(Y|A = 1) = \frac{1}{n} \sum_{i=1}^n \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_1 + \hat{\gamma}_2 W_i)$

- $\hat{E}(Y|A = 0) = \frac{1}{n} \sum_{i=1}^n \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_2 W_i)$

- 3 Calculate $\hat{\theta} = \hat{E}(Y|A = 1) - \hat{E}(Y|A = 0)$

Problem Setting

- Despite extensive literature and recommendations by regulators such as FDA and EMA, it remains highly **underutilized**.

Problem Setting

- Despite extensive literature and recommendations by regulators such as FDA and EMA, it remains highly **underutilized**.
- Problematic:
 - Resulting analyses are **inefficient** by not fully exploiting the available information in the data,
 - thereby **forfeiting the opportunity to reduce the required sample size**.

Potential Obstacles Leading to Underutilization

- 1 Many covariate adjustment methods are **incompatible with 'standard' group sequential designs (GSDs)**.
 - GSDs can **reduce the length** of a Phase 3 trial.
 - An obstacle for realizing precision gains from covariate adjustment as GSDs are commonly used for efficiency and ethical reasons.

Potential Obstacles Leading to Underutilization

- 1 Many covariate adjustment methods are **incompatible with 'standard' group sequential designs (GSDs)**.
 - GSDs can **reduce the length** of a Phase 3 trial.
 - An obstacle for realizing precision gains from covariate adjustment as GSDs are commonly used for efficiency and ethical reasons.
- 2 The **uncertainty** at the design stage about the **amount of precision gain and corresponding sample size reduction**.
 - An incorrect projection of a covariate's prognostic value risks an over- or underpowered future trial.

Outline

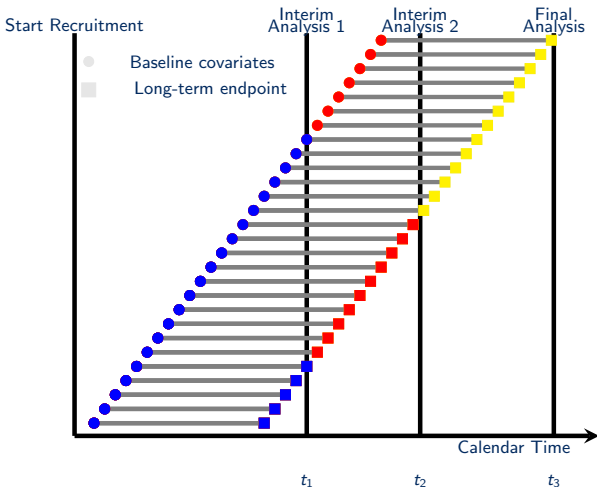
- 1 Background
- 2 Proposal: Combining Covariate Adjustment and GSDs**
- 3 Proposal: Combining Covariate Adjustment and Information-Adaptive Designs
- 4 Simulation Study
- 5 Discussion

Group Sequential Designs

- Entail analyzing the data at K different times t_1, \dots, t_K .

Group Sequential Designs

- Entail analyzing the data at K different times t_1, \dots, t_K .



Standardization Estimator at Interim Analyses

- Primary endpoint: binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: Standardization/G-computation

Standardization Estimator at Interim Analyses

- Primary endpoint: binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: Standardization/G-computation

At time t_k :

Standardization Estimator at Interim Analyses

- Primary endpoint: binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: Standardization/G-computation

At time t_k :

- 1 Fit logistic regression model for

$$P(Y = 1|A, W) = \text{logit}^{-1}(\gamma_0 + \gamma_1 A + \gamma_2 W),$$

in participants with complete follow up.

Standardization Estimator at Interim Analyses

- Primary endpoint: binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: Standardization/G-computation

At time t_k :

- 1 Fit logistic regression model for

$$P(Y = 1|A, W) = \text{logit}^{-1}(\gamma_0 + \gamma_1 A + \gamma_2 W),$$

in participants with complete follow up.

- 2 Compute standardized estimators for treatment specific means

- $\hat{E}_{t_k}(Y|A = 1) = \frac{1}{n'} \sum_{i=1}^{n'} \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_1 + \hat{\gamma}_2 W_i)$

- $\hat{E}_{t_k}(Y|A = 0) = \frac{1}{n'} \sum_{i=1}^{n'} \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_2 W_i)$

in all n' recruited patients.

Standardization Estimator at Interim Analyses

- Primary endpoint: binary.
- Estimand: $\theta = E(Y|A = 1) - E(Y|A = 0)$.
- Estimator: Standardization/G-computation

At time t_k :

- 1 Fit logistic regression model for

$$P(Y = 1|A, W) = \text{logit}^{-1}(\gamma_0 + \gamma_1 A + \gamma_2 W),$$

in participants with complete follow up.

- 2 Compute standardized estimators for treatment specific means

- $\hat{E}_{t_k}(Y|A = 1) = \frac{1}{n'} \sum_{i=1}^{n'} \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_1 + \hat{\gamma}_2 W_i)$

- $\hat{E}_{t_k}(Y|A = 0) = \frac{1}{n'} \sum_{i=1}^{n'} \text{logit}^{-1}(\hat{\gamma}_0 + \hat{\gamma}_2 W_i)$

in all n' recruited patients.

- 3 Calculate $\hat{\theta}_{t_k} = \hat{E}_{t_k}(Y|A = 1) - \hat{E}_{t_k}(Y|A = 0)$

Group Sequential Designs

- Entail analyzing the data at K different times t_1, \dots, t_K .

Group Sequential Designs

- Entail analyzing the data at K different times t_1, \dots, t_K .
- At each analysis time t_k :
 - Calculate an estimate $\hat{\theta}_{t_k}$.
 - Calculate a standardized test statistic $Z_k = Z(t_k) = \frac{\hat{\theta}_{t_k} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_k})}$.
 - Compare the Z_k to some critical value for that analysis.
 - Allow to stop early for efficacy and/or futility.

Group Sequential Designs

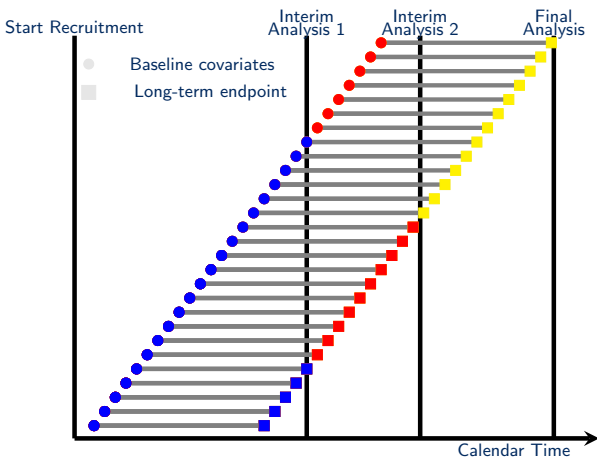
- Entail analyzing the data at K different times t_1, \dots, t_K .
- At each analysis time t_k :
 - Calculate an estimate $\hat{\theta}_{t_k}$.
 - Calculate a standardized test statistic $Z_k = Z(t_k) = \frac{\hat{\theta}_{t_k} - \theta_0}{\widehat{se}(\hat{\theta}_{t_k})}$.
 - Compare the Z_k to some critical value for that analysis.
 - Allow to stop early for efficacy and/or futility.
- Multiple looks at accumulating data **increase type I error**
 - Range of methods for defining the critical values for interim analyses.
(Pocock, 1977; O'Brien and Fleming, 1979; Lan and DeMets, 1983)

Group Sequential Designs: Independent Increments

- **Independent increments property:** $\hat{\theta}_{t_k}$ being asymptotically independent of all previous increments $\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}}$ for all $k' < k$.

Group Sequential Designs: Independent Increments

- **Independent increments property:** $\hat{\theta}_{t_k}$ being asymptotically independent of all previous increments $\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}}$, for all $k' < k$.



Group Sequential Designs: Independent Increments

- This property **holds for efficient estimators**.
 - ANCOVA with correctly specified model
 - G-computation and TMLE (if working models are correctly specified)
 - ...

Group Sequential Designs: Independent Increments

- This property **holds for efficient estimators**.

- ANCOVA with correctly specified model
- G-computation and TMLE (if working models are correctly specified)
- ...

- Unfortunately, a sequence of RAL estimators $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_K})$ does not necessarily have this property.

- e.g., G-computation and TMLE estimators when working models are misspecified

(e.g., Scharfstein et al., 1997; Jennison and Turnbull, 1997; Kim and Tsiatis, 2020; Rosenblum et al., 2015; Shoben and Emerson, 2014)

Group Sequential Designs: Independent Increments

- This property **holds for efficient estimators**.

- ANCOVA with correctly specified model
- G-computation and TMLE (if working models are correctly specified)
- ...

- Unfortunately, a sequence of RAL estimators $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_K})$ does not necessarily have this property.

- e.g., G-computation and TMLE estimators when working models are misspecified

(e.g., Scharfstein et al., 1997; Jennison and Turnbull, 1997; Kim and Tsiatis, 2020; Rosenblum et al., 2015; Shoben and Emerson, 2014)

- Proposal: **modifying** any RAL estimator so that it has the **independent increments property**.

Proposal: Motivation

- **Goal:** Obtain at each analysis time t_k an estimator $\tilde{\theta}_{t_k}$ that
 - 1 is consistent for θ ,
 - 2 is asymptotically linear,
 - 3 is asymptotically normal,
 - 4 is asymptotically as or more precise as the original estimator $\hat{\theta}_{t_k}$, and
 - 5 has the independent increments property.

Proposal: Motivation

- **Goal:** Obtain at each analysis time t_k an estimator $\tilde{\theta}_{t_k}$ that
 - 1 is consistent for θ ,
 - 2 is asymptotically linear,
 - 3 is asymptotically normal,
 - 4 is asymptotically as or more precise as the original estimator $\hat{\theta}_{t_k}$, and
 - 5 has the independent increments property.
- We will focus on finding the **linear combination**

$$\hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \lambda_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})$$

with **minimal variance**.

Proposal

- At $k = 1$, we let $\tilde{\theta}_{t_1} = \hat{\theta}_{t_1}$ and $\tilde{Z}_1 = Z_1 = \frac{\hat{\theta}_{t_1} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_1})}$.

Proposal

- At $k = 1$, we let $\tilde{\theta}_{t_1} = \hat{\theta}_{t_1}$ and $\tilde{Z}_1 = Z_1 = \frac{\hat{\theta}_{t_1} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_1})}$.
- At each subsequent analysis $k \geq 2$:
 - 1 Calculate $\hat{\theta}_{t_k}$ and estimate the covariance matrix of $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k})$.

Proposal

■ At $k = 1$, we let $\tilde{\theta}_{t_1} = \hat{\theta}_{t_1}$ and $\tilde{Z}_1 = Z_1 = \frac{\hat{\theta}_{t_1} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_1})}$.

■ At each subsequent analysis $k \geq 2$:

1 Calculate $\hat{\theta}_{t_k}$ and estimate the covariance matrix of $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k})$.

2 Solve

$$\left(\hat{\lambda}_1^{(k)}, \dots, \hat{\lambda}_{k-1}^{(k)}\right) = \arg \min_{(\lambda_1^{(k)}, \dots, \lambda_{k-1}^{(k)}) \in \mathbb{R}^{k-1}} \widehat{\text{Var}}\left\{\hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \lambda_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})\right\},$$

Proposal

■ At $k = 1$, we let $\tilde{\theta}_{t_1} = \hat{\theta}_{t_1}$ and $\tilde{Z}_1 = Z_1 = \frac{\hat{\theta}_{t_1} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_1})}$.

■ At each subsequent analysis $k \geq 2$:

1 Calculate $\hat{\theta}_{t_k}$ and estimate the covariance matrix of $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k})$.

2 Solve

$$\left(\widehat{\lambda}_1^{(k)}, \dots, \widehat{\lambda}_{k-1}^{(k)}\right) = \arg \min_{(\lambda_1^{(k)}, \dots, \lambda_{k-1}^{(k)}) \in \mathbb{R}^{k-1}} \widehat{\text{Var}}\left\{\hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \lambda_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})\right\},$$

3 Calculate $\tilde{\theta}_{t_k} = \hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \widehat{\lambda}_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})$, with

$$\widehat{\lambda}^{(k)} = \left\{ \widehat{\text{Var}}\left((\hat{\theta}_{t_k} - \hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k} - \hat{\theta}_{t_{k-1}})^t\right) \right\}^{-1} \cdot \widehat{\text{Cov}}\left(\hat{\theta}_{t_k}, (\hat{\theta}_{t_k} - \hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k} - \hat{\theta}_{t_{k-1}})^t\right)$$

Proposal

■ At $k = 1$, we let $\tilde{\theta}_{t_1} = \hat{\theta}_{t_1}$ and $\tilde{Z}_1 = Z_1 = \frac{\hat{\theta}_{t_1} - \theta_0}{\widehat{\text{se}}(\hat{\theta}_{t_1})}$.

■ At each subsequent analysis $k \geq 2$:

1 Calculate $\hat{\theta}_{t_k}$ and estimate the covariance matrix of $(\hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k})$.

2 Solve

$$\left(\widehat{\lambda}_1^{(k)}, \dots, \widehat{\lambda}_{k-1}^{(k)}\right) = \arg \min_{(\lambda_1^{(k)}, \dots, \lambda_{k-1}^{(k)}) \in \mathbb{R}^{k-1}} \widehat{\text{Var}}\left\{\hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \lambda_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})\right\},$$

3 Calculate $\tilde{\theta}_{t_k} = \hat{\theta}_{t_k} - \sum_{k'=1}^{k-1} \widehat{\lambda}_{k'}^{(k)} (\hat{\theta}_{t_k} - \hat{\theta}_{t_{k'}})$, with

$$\widehat{\lambda}^{(k)} = \left\{ \widehat{\text{Var}}\left((\hat{\theta}_{t_k} - \hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k} - \hat{\theta}_{t_{k-1}})^t\right) \right\}^{-1} \cdot \widehat{\text{Cov}}\left(\hat{\theta}_{t_k}, (\hat{\theta}_{t_k} - \hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k} - \hat{\theta}_{t_{k-1}})^t\right)$$

4 Calculate $\tilde{Z}_k = \frac{\tilde{\theta}_{t_k} - \theta_0}{\widehat{\text{se}}(\tilde{\theta}_{t_k})}$.

Outline

- 1 Background
- 2 Proposal: Combining Covariate Adjustment and GSDs
- 3 Proposal: Combining Covariate Adjustment and Information-Adaptive Designs**
- 4 Simulation Study
- 5 Discussion

Algorithm for Analysis Timing: Design Stage

- Specify the operating characteristics of the study

Algorithm for Analysis Timing: Design Stage

- Specify the operating characteristics of the study
- We **compute the maximum/total information** needed to preserve these operational characteristics

$$\left(\frac{z_{\alpha/2} + z_{\beta}}{\theta_A - \theta_0} \right)^2,$$

for a fixed design (no interim analyses), and

$$\left(\frac{z_{\alpha/2} + z_{\beta}}{\theta_A - \theta_0} \right)^2 IF$$

when data is sequentially monitored with the possibility of early stopping.

(Mehta and Tsiatis, 2001)

Algorithm for Analysis Timing: Information

- We propose to **monitor the accrued information**, $(\widehat{\text{se}}(\hat{\theta}_t))^{-2}$, through time t .

Algorithm for Analysis Timing: Information

- We propose to **monitor the accrued information**, $(\widehat{se}(\hat{\theta}_t))^{-2}$, through time t .
- We consider a trial with an interim analysis when 50% of the information is available:
 - We conduct the interim analysis at time t_1 when

$$(\widehat{se}(\hat{\theta}_{t_1}))^{-2} \geq 0.5 \cdot \left(\frac{z_{\alpha/2} + z_{\beta}}{\theta_A - \theta_0} \right)^2 IF.$$

- We conduct the final analysis at time t_2 when

$$(\widehat{se}(\hat{\theta}_{t_2}))^{-2} \geq \left(\frac{z_{\alpha/2} + z_{\beta}}{\theta_A - \theta_0} \right)^2 IF.$$

Outline

- 1 Background
- 2 Proposal: Combining Covariate Adjustment and GSDs
- 3 Proposal: Combining Covariate Adjustment and Information-Adaptive Designs
- 4 Simulation Study**
- 5 Discussion

MISTIE III trial (Stroke)

- Functional outcome: proportion of patients who achieved a modified Rankin Scale score of 0-3 at 365 days (**binary**).
- Estimand of interest: **risk difference**.
- Total sample size of approximately 498 patients (in original trial):

MISTIE III trial (Stroke)

- Functional outcome: proportion of patients who achieved a modified Rankin Scale score of 0-3 at 365 days (**binary**).
- Estimand of interest: **risk difference**.
- Total sample size of approximately 498 patients (in original trial):
 - 1:1 randomization
 - Power of 88% to detect an average effect size of 13% at a 5% significance level
 - Success rate: 25% in standard medical care group versus 38% in MISTIE group

MISTIE III trial (Stroke)

- Functional outcome: proportion of patients who achieved a modified Rankin Scale score of 0-3 at 365 days (**binary**).
- Estimand of interest: **risk difference**.
- Total sample size of approximately 498 patients (in original trial):
 - 1:1 randomization
 - Power of 88% to detect an average effect size of 13% at a 5% significance level
 - Success rate: 25% in standard medical care group versus 38% in MISTIE group
- We will focus on information instead of sample size!

Simulation Study: $K = 2$

- We perform interim analysis when 50% of the (total) information is available
 - Total information: 648

		$\theta = 0.13$ (Alternative)		
		Power	ASN	AAT
Original estimators $\hat{\theta}_{t_k}$	Unadjusted	88.3%	534	1566
	Standardization	87.1%	431	1299
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization	87.0%	431	1299

ASN: average sample number; AAT: average analysis time (days).

Note: We did a small sample size correction for standardization estimator.

Conclusion under alternative:

19% reduction of sample size due to covariate adjustment

Simulation Study: $K = 2$

- We perform interim analysis when 50% of the (total) information is available
 - Total information: 648

		$\theta = 0$ (Null)		
		Type I	ASN	AAT
Original estimators $\hat{\theta}_{t_k}$	Unadjusted	5.29%	628	2014
	Standardization	5.06%	449	1542
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization	5.05%	449	1542

AAT: average analysis time (days); ASN: average sample number.

Note: We did a small sample size correction for standardization estimator.

Conclusion under null:

29% reduction of sample size due to covariate adjustment

Outline

- 1 Background
- 2 Proposal: Combining Covariate Adjustment and GSDs
- 3 Proposal: Combining Covariate Adjustment and Information-Adaptive Designs
- 4 Simulation Study
- 5 Discussion**

Discussion

- We performed additional simulations under violation of independent increments property:

			Type I
Original estimators $\hat{\theta}_{t_k}$	Standardization		5.37%
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization		5.07%

Discussion

- We performed additional simulations under violation of independent increments property:

			Type I
Original estimators $\hat{\theta}_{t_k}$	Standardization		5.37%
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization		5.07%

- Simulations have only shown small deviations from independent increment structure.

Discussion

- We performed additional simulations under violation of independent increments property:

			Type I
Original estimators $\hat{\theta}_{t_k}$	Standardization		5.37%
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization		5.07%

- Simulations have only shown small deviations from independent increment structure.
- In practice, underlying **data-generating mechanism is unknown**.

Discussion

- We performed additional simulations under violation of independent increments property:

			Type I
Original estimators $\hat{\theta}_{t_k}$	Standardization		5.37%
Orthogonalized estimators $\tilde{\theta}_{t_k}$	Standardization		5.07%

- Simulations have only shown small deviations from independent increment structure.
- In practice, underlying **data-generating mechanism is unknown**.
- **Safer** to use the proposal as it guarantees to maintain the Type I error in large samples.

Discussion

Importantly, works for **all kind of endpoints and estimands** as long as the considered estimators are **consistent for θ and asymptotically linear**.

(Not necessarily covariate adjusted estimators!).



Thank you for your attention!

Interested? <https://doi.org/10.48550/arXiv.2201.12921>

E-mail: kelly.vanlancker@ugent.be

The opinions in this presentation are of the author and do not necessarily represent those of anyone else.

References I

- Benkeser, D., I. Díaz, A. Luedtke, J. Segal, D. Scharfstein, and M. Rosenblum (2020). Improving precision and power in randomized trials for covid-19 treatments using covariate adjustment, for binary, ordinal, and time-to-event outcomes. *Biometrics*.
- Jennison, C. and B. W. Turnbull (1997). Group-sequential analysis incorporating covariate information. *Journal of the American Statistical Association* 92(440), 1330–1341.
- Jiang, F., L. Tian, H. Fu, T. Hasegawa, and L. J. Wei (2018). Robust alternatives to ANCOVA for estimating the treatment effect via a randomized comparative study. *Journal of the American Statistical Association* 0, 1–37.
- Kim, K. and A. A. Tsiatis (2020). Independent increments in group sequential tests: a review. *SORT-Statistics and Operations Research Transactions*, 223–264.

References II

- Koch, G. G., C. M. Tangen, J.-W. Jung, and I. A. Amara (1998). Issues for covariance analysis of dichotomous and ordered categorical data from randomized clinical trials and non-parametric strategies for addressing them. *Stat. Med.* 17(15-16), 1863–1892.
- Lan, G. K. and D. L. DeMets (1983). Discrete sequential boundaries for clinical trials. *Biometrika* 70(3), 659–663.
- Mehta, C. R. and A. A. Tsiatis (2001). Flexible sample size considerations using information-based interim monitoring. *Drug information journal: DIJ/Drug Information Association* 35(4), 1095–1112.
- Moore, K. and M. J. van der Laan (2009a). Covariate adjustment in randomized trials with binary outcomes: Targeted maximum likelihood estimation. *Stat. Med.* 28(1), 39–64.

References III

- Moore, K. L. and M. J. van der Laan (2009b). Increasing power in randomized trials with right censored outcomes through covariate adjustment. *Journal of Biopharmaceutical Statistics* 19(6), 1099–1131. PMID: 20183467.
- O'Brien, P. C. and T. R. Fleming (1979). A multiple testing procedure for clinical trials. *Biometrics*, 549–556.
- Pocock, S. J. (1977). Group sequential methods in the design and analysis of clinical trials. *Biometrika* 64(2), 191–199.
- Rosenblum, M., T. Qian, Y. Du, , and H. Qiu (2015). Adaptive enrichment designs for randomized trials with delayed endpoints, using locally efficient estimators to improve precision. *Johns Hopkins University, Dept. of Biostatistics Working Papers*. <https://biostats.bepress.com/jhubiostat/paper275>.

References IV

- Rubin, D. and M. van der Laan (2008). Covariate adjustment for the intention-to-treat parameter with empirical efficiency maximization. *U.C. Berkeley Division of Biostatistics Working Paper Series. Working Paper 229*, <https://biostats.bepress.com/ucbbiostat/paper229>.
- Scharfstein, D. O., A. A. Tsiatis, and J. M. Robins (1997). Semiparametric efficiency and its implication on the design and analysis of group-sequential studies. *J Am Stat Assoc* 92(440), 1342–1350.
- Shoben, A. B. and S. S. Emerson (2014). Violations of the independent increment assumption when using generalized estimating equation in longitudinal group sequential trials. *Statistics in medicine* 33(29), 5041–5056.

References V

- Tsiatis, A. A., M. Davidian, M. Zhang, and X. Lu (2008). Covariate adjustment for two-sample treatment comparisons in randomized clinical trials: a principled yet flexible approach. *Statistics in medicine* 27(23), 4658–4677.
- Yang, L. and A. Tsiatis (2001). Efficiency study of estimators for a treatment effect in a pretest-posttest trial. *The American Statistician* 55(4), 314–321.
- Zhang, D. (2009). Lecture notes for statistical principles of clinical trials (modified from dr. a. tsiatis' lecture notes).
- Zhang, M. (2015, Jan). Robust methods to improve efficiency and reduce bias in estimating survival curves in randomized clinical trials. *Lifetime Data Analysis* 21(1), 119–137.

Proposal: Variance

- Estimate the variance of $\tilde{\theta}_{t_k}$ as

$$\widehat{\text{se}}(\tilde{\theta}_k)^2 = (-\widehat{\lambda}^{(k)})^t, 1) \widehat{\text{Cov}} \left((\hat{\theta}_{t_k} - \hat{\theta}_{t_1}, \dots, \hat{\theta}_{t_k} - \hat{\theta}_{t_{k-1}}, \hat{\theta}_{t_k})^t \right) (-\widehat{\lambda}^{(k)})^t, 1)^t.$$

- $n \cdot \widehat{\text{se}}(\tilde{\theta}_k)^2$ is a **consistent** estimate for the asymptotic variance $n \cdot \text{Var}(\tilde{\theta}_{t_k})$.
- This guarantees **asymptotically correct hypothesis testing and confidence intervals**.

Algorithm for Analysis Timing: (Dis)advantages

- The **information-adaptive design** is well suited for being adopted for covariate adjusted estimators:
 - We do **not have to prespecify the prognostic value of the covariates** nor other nuisance parameters.
 - When the estimator is more efficient than unadjusted estimator, covariate adjustment can lead to a **shorter trial** due to faster information accrual.

Algorithm for Analysis Timing: (Dis)advantages

- The **information-adaptive design** is well suited for being adopted for covariate adjusted estimators:
 - We do **not have to prespecify the prognostic value of the covariates** nor other nuisance parameters.
 - When the estimator is more efficient than unadjusted estimator, covariate adjustment can lead to a **shorter trial** due to faster information accrual.
- **Administrative inconvenience**: it does not give an idea to the investigators about the necessary resources (i.e., length of study, sample size, ...).
 - We suggest to posit some **guesses on the nuisance parameters**, and
 - use the **emerging data to evaluate** whether the maximum information can be reached with the available resources.